### The Local Limit Theorem and the Almost Sure Local Limit Theorem

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Department of Mathematics "L. Tonelli" University of Pisa ITALY Local Limit Theorems (LLT)

- The DeMoivre–Laplace theorem
- Gnedenko's Theorem
- Necessary and sufficient conditions for the LLT
- Necessary and sufficient conditions for the LLT with rate
- The general LLT for random variables in the domain of attraction of a stable distribution

### Almost Sure Local Limit Theorems (ASLLT)

- Motivation: → Almost Sure Central Limit Theorem (ASCLT)
- ASLLT for random sequences in the domain of attraction of the normal law
  - ASLLT with rate (random sequences with moment  $2 + \epsilon$ )
  - ASLLT for random sequences with second moment
- ASLLT for random sequences in the domain of attraction of a stable law with lpha < 2 (without second moment)

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### The DeMoivre–Laplace theorem

### Theorem

Let  $g_n(k)$  be the probability of getting k heads in n tosses of a coin which gives a head with probability p. Then

$$\lim_{n\to\infty}\frac{g_n(k)}{\left(\frac{1}{\sqrt{2npq}}e^{-\frac{(k-np)^2}{2npq}}\right)}=1,$$

uniformly for k such that  $\left|\frac{k-np}{\sqrt{npq}}\right|$  remains bounded.

- Abraham DeMoivre proved it only for a fair coin (p = 1/2) in Approximatio ad Summam Terminorum Binomii  $(a + b)^n$  in Seriem expansi (1733).
- Pierre-Simon Laplace proved it in full generality in *Théorie* Analytique des probabilités (1812).

In the DeMoivre Theorem  $\longrightarrow p = E[X_1]$  and  $pq = VarX_1$ .

We could expect

#### Theorem

Let  $(X_n)_{n \ge 1}$  be a sequence of i.i.d. random variables, with  $E[X_1] = \mu$ ,  $VarX_1 = \sigma^2$ . Then

$$P(S_n = k) \approx rac{1}{\sqrt{2\pi n}\sigma} e^{-rac{(k-n\mu)^2}{2n\sigma^2}}$$

if  $\left|\frac{k-n\mu}{\sqrt{npq}}\right|$  is bounded.

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But certainly this cannot be true in general: if

$$X_1 = egin{cases} -1 & ext{with probability } rac{1}{2} \ 1 & ext{with probability } rac{1}{2} \end{cases}$$

If k is odd, then  $P(S_{2n} = k) = 0$ .

If k is even, say k = 2h, we should obtain (with 2n in place of n and 2h in place of k,  $\mu = 0$ ,  $\sigma^2 = 1$ ).

$$P(S_{2n}=2h)\approx\frac{1}{2\sqrt{\pi n}}e^{-\frac{h^2}{n}},$$

with  $\frac{h}{\sqrt{n}}$  bounded.

### In particular

$$\longrightarrow$$
 for  $(h_n)$  such that  $\frac{h_n}{\sqrt{n}} \rightarrow_n \frac{x}{\sqrt{2}}$ :  
 $P(S_{2n} = 2h_n) \approx \frac{1}{2\sqrt{\pi n}} e^{-\frac{x^2}{2}}.$ 

#### On the contrary

## Theorem If $\frac{h_n}{\sqrt{n}} \to_n \frac{x}{\sqrt{2}}$ $\lim_{n \to \infty} \sqrt{n} P(S_n = 2h_n) = \frac{1}{\sqrt{\pi}} e^{-\frac{x^2}{2}}.$ (2)

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Hence a general theorem should roughly state that

$$P(S_n = k) \approx \frac{c}{\sqrt{2\pi n\sigma}} e^{-\frac{(k-n\mu)^2}{2n\sigma^2}}$$

if k is a value of  $S_n$  and  $\left|\frac{k-n\mu}{\sqrt{npq}}\right|$  is bounded, for a suitable c. But what is c? Comparing (1) and (2), we notice that the main difference is a factor of 2 in the second member of (2); where does it come from? We also notice that in this case the support of  $S_{2n}$  is concentrated on even integers, and two successive even integers differ by 2. So one guesses that c = 2 in our case, and in general c is maybe connected with the gap between successive values of  $S_n$ .

### Preliminaries

$$\mathcal{L}(a,\lambda) := a + \lambda \mathbb{Z} = \{a + \lambda k, k \in \mathbb{Z}\}.$$

### Definition

A random variable X has a **lattice distribution** if there exist two constant a and  $\lambda > 0$  such that  $P(X \in \mathcal{L}(a, \lambda)) = 1$ .

 $\phi = \mathbb{E}[e^{itX}] =$  the characteristic function of X.

Link between lattice distribution and the behaviour  $\phi$  :

#### Theorem

There are only three possibilities:

(i) there exists a  $t_0 > 0$  such that  $|\phi(t_0)| = 1$  and  $|\phi(t)| < 1$  for every  $0 < t < t_0$ . In this case X has a lattice distribution. (ii)  $|\phi(t)| < 1$  for every  $t \neq 0$  (non lattice distribution). (iii)  $|\phi(t)| = 1$  for every  $t \in \mathbb{R}$ : In this case X is constant a.s. (degenerate distribution). The proof shows in particular that

### Corollary

In case (i) of the preceding Theorem, we have

$$rac{2\pi}{t_0}= ext{max}\{\lambda>0:\exists\, a\in\mathbb{R},\ extsf{P}(X\in\mathcal{L}(a,\lambda))=1\}.$$

Hence

### Definition

In case (i) of the preceding Theorem, the number

$$\Lambda = \frac{2\pi}{t_0} = \max\{\lambda > 0 : \exists a \in \mathbb{R}, P(X \in \mathcal{L}(a, \lambda)) = 1\}$$

is called the (maximal) span of the distribution of X.

• (i) Let

$$X_1 = egin{cases} -1 & ext{with probability}\,rac{1}{2} \ 1 & ext{with probability}\,rac{1}{2}. \end{cases}$$

Then  $\phi(t) = \frac{1}{2}(e^{it} + e^{-it}) = \cos t$ , and  $|\phi(t)| = 1$  if and only if  $t = n\pi$ ,  $n \in \mathbb{Z}$ . Hence  $t_0 = \pi$ , and the maximal span of the distribution is  $\frac{2\pi}{t_0} = 2$ .

- (ii) Let  $X_1$  have standard gaussian law. Then  $\phi(t) = e^{-\frac{t^2}{2}}$ . In this case  $|\phi(t)| = 1$  only for t = 0.
- (iii) If  $X_1 = c$  (*c* some constant) we have  $\phi(t) = e^{itc}$ , and  $|\phi(t)| = 1$  for every *t*.

 $(X_n)_{n \ge 1}$  sequence of i.i.d random variables,  $\mathbf{E}[X_i] = \mu$ ,  $\mathbf{Var}X_i = \sigma^2$  (finite) with lattice distribution.  $\Lambda = \text{maximal span}$ .

$$S_n = X_1 + \cdots + X_n.$$

$$P(X_i \in \mathcal{L}(a, \Lambda)) = 1 \Longrightarrow P(S_n \in \mathcal{L}(na, \Lambda) = 1.$$

### Theorem

We have

$$\lim_{n\to\infty}\sup_{N\in\mathcal{L}(na,\Lambda)}\left|\frac{\sqrt{n}}{\Lambda}P(S_n=N)-\frac{1}{\sqrt{2\pi}\sigma}e^{-\frac{(N-n\mu)^2}{2n\sigma^2}}\right|=0.$$

### Gnedenko's Local Limit Theorem (1948)

Some heuristics ( $\mu = 0$ )

By the CLT

$$\begin{split} & P\Big(S_n = N\Big) \approx P\Big(N - \frac{\Lambda}{2} \leqslant S_n \leqslant N + \frac{\Lambda}{2}\Big) \\ &= P\Big(\frac{N}{\sqrt{n\sigma}} - \frac{\Lambda}{2\sigma\sqrt{n}} \leqslant \frac{S_n}{\sigma\sqrt{n}} \leqslant \frac{N}{\sqrt{n\sigma}} - \frac{\Lambda}{2\sigma\sqrt{n}}\Big) \\ &\approx \int_{\frac{N}{\sqrt{n\sigma}} - \frac{\Lambda}{2\sigma\sqrt{n}}}^{\frac{N}{\sqrt{n\sigma}} + \frac{\Lambda}{2\sigma\sqrt{n}}} \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt = \int_{\frac{N}{\sqrt{n}} - \frac{\Lambda}{2\sqrt{n}}}^{\frac{N}{\sqrt{n}} + \frac{\Lambda}{2\sigma\sqrt{n}}} \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{t^2}{2\sigma^2}} dt \\ &\approx \frac{\Lambda}{\sqrt{n}} \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{N^2}{2n\sigma^2}}. \end{split}$$

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# Gnedenko's Local Limit Theorem (necessary and sufficient conditions for the LLT)

Actually, the complete formulation of Gnedenko's result is

#### Theorem

With the same assumptions as above, in order that

$$\lim_{n\to\infty}\sup_{N\in\mathcal{L}(na,\lambda)}\left|\frac{\sqrt{n}}{\lambda}P(S_n=N)-\frac{1}{\sqrt{2\pi}\sigma}e^{-\frac{(N-n\mu)^2}{2n\sigma^2}}\right|=0,$$

it is necessary and sufficient that  $\lambda = \Lambda$ .

### Necessary and sufficient conditions for the Local Limit Theorem with rate

The following result completes the theory

### Theorem

With the same assumptions as in Theorem 8, in order that

$$\sup_{N\in\mathcal{L}(na,\lambda)}\left|\frac{\sqrt{n}}{\lambda}P(S_n=N)-\frac{1}{\sqrt{2\pi\sigma}}e^{-\frac{(N-n\mu)^2}{2n\sigma^2}}\right|=O(n^{-\alpha})$$

with

$$0 < \alpha < \frac{1}{2}$$

*it is necessary and sufficient that the following conditions are satisfied* 

 $(i)\lambda = \Lambda;$  (ii) if F denotes the distribution function of  $X_1$ , then, as  $u \to \infty$ ,  $\int_{|x|>u} x^2 F(dx) = O(u^{-2\alpha}).$   $(X_n)_{n\geq 1}$  sequence of i.i.d. random variables .  $\phi =$  characteristic function with  $|\phi(t)| < 1$  for every  $t \neq 0$ .

#### Remark

*In the nonlattice case, most characteristic functions verify the Cramer's condition, i.e.* 

 $\limsup_{t o \infty} |\phi(t)| < 1.$ 

### Remark

There exist characteristic functions of nonlattice random variables, that do not verify Cramer's condition.

An example is

$$\phi(t) = \prod_{k=1}^{\infty} \cos\left(\frac{t}{k!}\right).$$

 $|\phi(t)| = 1 \iff \frac{t}{k!}$  is a multiple of  $\pi$  for each integer k: impossible unless t = 0.

But

$$1-\phi(2\pi N!) \to 0, \qquad N \to \infty.$$

### The following result holds

#### Theorem

Let  $(X_n)_{n \ge 1}$  be sequence of i.i.d. nonlattice random variables, with  $\mathbf{E}[X_1] = \mu$ ,  $\mathbf{Var}X_1 = \sigma^2 < \infty$ . If  $\frac{x_n}{\sqrt{n}} \to x$  and a < b,  $\lim_{n \to \infty} \sqrt{n}P(S_n - n\mu \in (x_n + a, x_n + b)) = (b - a)\frac{1}{\sqrt{2\pi}\sigma}e^{-\frac{x^2}{2\sigma^2}}$ .

The heuristics are as before.

With some further properties on  $|\phi|$ :

### Theorem

If  $|\phi|$  is integrable, then  $\frac{S_n - n\mu}{\sigma\sqrt{n}}$  has a density  $f_n$ ; moreover  $f_n$  tends undiformly to the standard normal density

$$\eta(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

The general Local Limit Theorem for random variables in the domain of attraction of a stable distribution: preliminaries

 $(X_n)_{n \ge 1}$  i.i.d. with (common) distribution F (not necessarily lattice) and partial sums  $S_n = X_1 + \cdots + X_n$ . G = distribution.

### Definition

The **domain of attraction** of *G* is the set of distributions *F* having the following property: there exists two sequences  $(a_n)$  and  $(b_n)$  of real numbers, with  $b_n \rightarrow_n \infty$ , such that

$$\frac{S_n - a_n}{b_n} \stackrel{\mathcal{L}}{\longrightarrow} G$$

as  $n \to \infty$ .

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The general Local Limit Theorem for random variables in the domain of attraction of a stable distribution: stable distributions

G possesses a domain of attraction iff G is *stable*, i.e.

### Definition

A non-degenerate distribution G is **stable** if it satisfies the following property:

let  $X_1$  and  $X_2$  be independent variables with distribution G; for any constants a > 0 and b > 0 the random variable  $aX_1 + bX_2$  has the same distribution as  $cX_1 + d$  for some constants c > 0 and d.

The general Local Limit Theorem for random variables in the domain of attraction of a stable distribution: alternative definition of stable distributions

### Definition

G is stable if its characteristic function can be written as

$$arphi(t;\mu,c,lpha,eta) = \exp\left[ it\mu \!-\! |ct|^lpha \left(1\!-\!ieta\,\mathrm{sgn}(t)\Phi
ight) 
ight]$$

where  $\alpha \in (0,2]$ ,  $\mu \in \mathbb{R}$ ,  $\beta \in [-1,1]$ ; sgn(t) is just the sign of t and

$$\Phi = \begin{cases} \tan \frac{\pi \alpha}{2} & \text{if } \alpha \neq 1 \\ -\frac{2}{\pi} \log |t| & \text{if } \alpha = 1. \end{cases}$$

The parameter  $\alpha$  is the **exponent** of the distribution.

### Remark

The normal law is stable with exponent  $\alpha = 2$ .

### The general Local Limit Theorem for random variables in the domain of attraction of a stable distribution

### Theorem

Let  $X_n$  have lattice distribution with maximal span  $\Lambda$ . In order that, for some choice of constants  $a_n$  and  $b_n$ 

$$\lim_{n\to\infty}\sup_{N\in\mathcal{L}(na,\lambda)}\left|\frac{b_n}{\lambda}P(S_n=N)-g\left(\frac{N-a_n}{b_n}\right)\right|=0,$$

where g is the density of some stable distribution G with exponent  $0 < \alpha \le 2$ ,

it is necessary and sufficient that

(i) the common distribution F of the  $X_n$  belongs to the domain of attraction of G;

(ii)  $\lambda = \Lambda$  (i.e. maximal).

Almost Sure Local Limit Theorems (ASLLT): the motivation (starting from the Almost Sure <u>Central</u> Limit Theorem)

$$(X_n)_{n\geq 1}$$
 i.i.d with  $\mathbf{E}[X_1] = \mu$ ,  $\mathbf{Var}X_1 = \sigma^2$ .

$$Z_n = \frac{S_n - n\mu}{\sigma\sqrt{n}}.$$

### **Central Limit Theorem**

## Theorem For every $x \in \mathbb{R}$ $\mathbf{E}[1_{\{Z_n \leq x\}}] = P(Z_n \leq x) \xrightarrow[n]{} \Phi(x) := \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt$

Almost Sure Local Limit Theorems (ASLLT): the motivation (starting from the Almost Sure <u>Central</u> Limit Theorem)

### Almost Sure Central Limit Theorem

## Theorem $P-a. \ s., \ for \ every \ x \in \mathbb{R}$ $\frac{1}{\log n} \sum_{h=1}^{n} \frac{1}{h} \underbrace{\mathbb{1}_{\{Z_h \leq x\}}}_{n} \longrightarrow \Phi(x).$

### By analogy (for the case of Gnedenko's Theorem)

$$\kappa_n \in \mathcal{L}(\mathit{na}, \Lambda)$$
 such that  $rac{\kappa_n - n\mu}{\sqrt{n}} 
ightarrow \kappa.$ 

Gnedenko's Theorem  $\Longrightarrow$ 

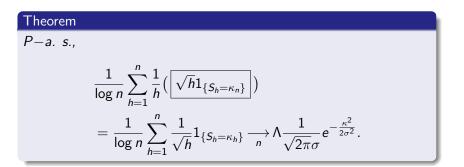
$$\mathbf{E}\left[\sqrt{n}\mathbf{1}_{\{S_n=\kappa_n\}}\right] = \sqrt{n}P(S_n=\kappa_n) \to \Lambda \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{\kappa^2}{2\sigma^2}}.$$

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### Comparing with the case of the Central Theorem:

### Tentative Almost Sure Local Limit Theorem



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### • In 1951 Chung and Erdös proved

### Theorem

Let  $(X_n)_{n \ge 1}$  be a centered Bernoulli process with parameter p. Then

$$\frac{1}{\log n}\sum_{h=1}^{n}\frac{1}{\sqrt{h}}\mathbf{1}_{\{S_{h}=0\}}\longrightarrow \frac{1}{\sqrt{2\pi p(1-p)}}, \qquad a.s.$$

This is a particular case of our tentative ASLLT: just take  $\kappa_n = np$ .

### • In 1993 Csáki, Földes and Révész proved

### Theorem

Let  $(X_n)_{n \ge 1}$  be i.i.d. centered and with finite third moment. Then

$$\frac{1}{\log n} \sum_{h=1}^{n} \frac{1}{p_h} \mathbb{1}_{\{a_h \leqslant S_h \leqslant b_h\}} \xrightarrow{n} \mathbb{1}, \qquad a.s.$$

where  $p_n = P(a_n \leq S_n \leq b_n)$ .

This generalizes the Chung-Erdös Theorem: just take  $a_n = b_n = 0$  and recall Gnedenko's Theorem.

# The ASLLT for random sequences in the domain of attraction of the normal law

 $(X_n)_{n \ge 1}$  i.i.d. having lattice distribution F with maximal span  $\Lambda$ ;  $\mathbf{E}[X_1] =: \mu$ ,  $\mathbf{Var}X_1 =: \sigma^2$  (finite). We assume  $\mu = 0$  and  $\sigma^2 = 1$ (no loss of generality).

### Definition

 $(X_n)_{n \ge 1}$  satisfies an ASLLT if

$$\lim_{n \to \infty} \frac{1}{\log n} \sum_{h=1}^{n} \frac{1}{\sqrt{h}} \mathbb{1}_{\{S_h = \kappa_h\}} \stackrel{a.s.}{=} \Lambda \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{\kappa^2}{2\sigma^2}}$$

for any sequence of integers  $(\kappa_n)_{n \ge 1}$  in  $\mathcal{L}(na, \Lambda)$  such that

$$\lim_{n\to\infty}\frac{\kappa_n-n\mu}{\sqrt{n}}=\kappa.$$

### Theorem

Let  $\epsilon > 0$  and assume that  $\mathbf{E}[|X_1^{2+\epsilon}|] < \infty$ . Then  $(X_n)_{n \ge 1}$  satisfies an ASLLT. Moreover, if the sequence  $(\kappa_n)_{n \ge 1}$  verifies the stronger condition

$$\frac{\kappa_n - n\mu}{\sqrt{n}} = \kappa + O_{\delta}((\log n)^{-1/2 + \delta})$$

then

$$\sum_{h=1}^{n} \frac{1}{\sqrt{h}} \mathbb{1}_{\{S_h = \kappa_h\}} = \Lambda \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{\kappa^2}{2\sigma^2}} + O_{\delta}((\log n)^{-1/2+\delta}).$$

### Remark

If  $\mathbf{E}[|X_1^{2+\epsilon}|] < \infty$  for some positive  $\epsilon$ , then the condition of Gnedenko's Theorem with rate, i.e.

$$\sup_{N \in \mathcal{L}(na,\lambda)} \left| \frac{\sqrt{n}}{\lambda} P(S_n = N) - \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(N-n\mu)^2}{2n\sigma^2}} \right| = O(n^{-\alpha}),$$
$$0 < \alpha < \frac{1}{2}$$

is satisfied with  $\alpha = \epsilon/2$ . In fact

$$\int_{|x|\geqslant u} x^2 F(dx) = \int_{|x|\geqslant u} |x|^{2+\epsilon} |x|^{-\epsilon} F(dx) \leqslant \mathbf{E}[|X_1^{2+\epsilon}|] u^{-\epsilon}.$$

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- (i) a suitable correlation inequality;
- (ii) Gnedenko's Theorem with rate;
- (iii) the notion of quasi orthogonal system.

### Correlation inequality

$$Y_h = \sqrt{h} \big( \mathbb{1}_{\{S_h = \kappa_h\}} - P(S_h = \kappa_h) \big).$$

### Proposition

Assume that

$$r(n) := \sup_{N \in \mathcal{L}(na,\Lambda)} \left| \frac{\sqrt{n}}{\Lambda} P(S_n = N) - \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(N-n\mu)^2}{2n\sigma^2}} \right| = O(n^{-\alpha}),$$
  
$$0 < \alpha < \frac{1}{2}$$

Then there exists a constant C such that, for all integers m, n with  $1 \le m < n$ 

$$\left|\operatorname{\mathbf{Cov}}(Y_m,Y_n)
ight|\leq C\Big(rac{1}{\sqrt{rac{n}{m}}-1}+\sqrt{rac{n}{n-m}}\cdotrac{1}{(n-m)^lpha}\Big).$$

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### The notion of quasi orthogonal system

Kac-Salem-Zygmund definition of quasi-orthogonal system

### Definition

A sequence of functions  $\Psi := (f_n)_{n \ge 1}$  defined on a Hilbert space  $\mathcal{H}$  is said **quasi-orthogonal** if the quadratic form on  $\ell^2$ :  $(x_n) \mapsto \sum_{h,k} \langle f_h, f_k \rangle x_h x_k$  is bounded (as a quadratic form).

### Weber's criterion for quasi-orthogonality

#### Lemma

In order that  $\Psi := (f_n)_{n \ge 1}$  be a quasi-orthogonal system, it is sufficient that

$$\sup_{h}\sum_{k}|\langle f_{h},f_{k}\rangle|<\infty.$$

### Remark

If  $\mathcal{H} = L^2(T)$ , where  $(T, \mathcal{A}, \mu)$  is some probability space, then  $\sum_{h,k} \langle f_h, f_k \rangle x_h x_k = \sum_{h,k} (\int f_h f_k d\mu) x_h x_k$ . By Rademacher–Menchov Theorem, it is seen that the series  $\sum_n c_n f_n$ converges if for instance  $c_n = n^{-\frac{1}{2}} (\log n)^{-b}$  with  $b > \frac{3}{2}$ . (i) Any  $\rho > 1$  fixed.

The basic correlation inequality  $\implies$   $Z_j = \sum_{\rho^j \leqslant h < \rho^{j+1}} \frac{Y_h}{h}$  is quasi-orthogonal. (ii) By the preceding remark

$$\sum_j \frac{Z_j}{\sqrt{j}(\log j)^b}$$
 converges as soon as  $b > \frac{3}{2}$ .

(iii) By Kronecker's Lemma

$$\frac{1}{\sqrt{n}(\log n)^b} \sum_{j=1}^n Z_j = \frac{1}{\sqrt{n}(\log n)^b} \sum_{j=1}^n \sum_{\rho^j \leqslant h < \rho^{j+1}} \frac{Y_h}{h}$$
$$= \frac{1}{\sqrt{n}(\log n)^b} \sum_{1 \leqslant h < \rho^{n+1}} \frac{Y_h}{h} \xrightarrow{} 0.$$

# Main steps of the proof

(iv) The preceding relation yields easily (details omitted)

$$\frac{\sqrt{\log t}}{(\log \log t)^b} \Big( \frac{1}{\log t} \sum_{h \leqslant t} \frac{Y_h}{h} \Big) = \frac{1}{\sqrt{\log t} (\log \log t)^b} \sum_{h \leqslant t} \frac{Y_h}{h} \xrightarrow{t} 0;$$
  
since  $\frac{\sqrt{\log t}}{(\log \log t)^b} \longrightarrow_t \infty$ , this implies

$$\frac{1}{\log t} \sum_{h \leqslant t} \frac{Y_h}{h} = \frac{1}{\log t} \sum_{h \leqslant t} \frac{\mathbb{1}_{\{S_h = \kappa_h\}}}{\sqrt{h}} - \frac{1}{\log t} \sum_{h \leqslant t} \frac{\sqrt{h}P(S_h = \kappa_h)}{h} \xrightarrow{t} 0$$

(v) Last, by Gnedenko's Theorem

$$\frac{1}{\log t} \sum_{h \leqslant t} \frac{\sqrt{h} P(S_h = \kappa_h)}{h} \xrightarrow{t} \Lambda \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{\kappa^2}{2\sigma^2}},$$

and the result follows.

The second part of the Theorem is proved similarly.

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## ASLLT for random sequences with second moment

When only the second moment is available, the ASLLT with rate doen't hold (see remark). M. Weber proved the ASLLT afresh:

with the additional assumption (v<sub>k</sub>, k ∈ Z are the elements of the lattice)

$$P(X = v_k) \wedge P(X = v_{k+1}) > 0$$
 for some  $k \in \mathbb{Z}$ . (3)

With (3), a basic correlation inequality holds, again for  $Y_h = \sqrt{h} (1_{\{S_h = \kappa_h\}} - P(S_h = \kappa_h))$ :

#### Proposition

Assume that (3) holds. Then there exists a constant C (depending on the sequence  $(\kappa_n)$ ) such that, for all integers m, n with  $1 \le m < n$ 

$$\left|\operatorname{Cov}(Y_m,Y_n)\right| \leq C\Big(rac{1}{\sqrt{rac{n}{m}}-1}+\sqrt{rac{n}{n-m}}\cdotrac{1}{n-m}\Big).$$

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in the general case

The Local Limit Theorem and the Almost Sure Local Limit Theo

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Main ingredients similar to the previous ones:

- the above correlation inequality;
- Gnedenko's Theorem (without rate);
- **(3)** the notion of **quasi orthogonal system.**

Recent work by R. Giuliano and Z. Szewczak (work in progress).

- (X<sub>n</sub>)<sub>n≥1</sub> sequence of i.i.d. random variables; common distribution F is in the domain of attraction of a stable distribution G (having density g) with exponent α (0 < α < 2).</li>
- $(a_n)$  and  $(b_n)$  as before;  $b_n = L(n)n^{1/\alpha}$ , L slowly varying in Karamata's sense.
- $X_1$  with lattice distribution;  $\Lambda = \max$  maximal span.
- F and G symmetric  $\implies a_n = 0$ .

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### IMPORTANT

Since  $X_1$  doesn't possess second moment, the discussion of the preceding section doesn't work.

But we can use similar ingredients as before:

- a correlation inequality;
- General LLT Theorem (for stable distributions);
- Gaal–Koksma LLN.

## The correlation inequality when $0 < \alpha < 2$

### Proposition

• For every pair (m, n) of integers, with  $1 \le m < n$ ,

$$egin{aligned} b_m b_n \Big| P(S_m = \kappa_m, S_n = \kappa_n) - P(S_m = \kappa_m) P(S_n = \kappa_n) \Big| \ &\leq C \Big\{ \Big( rac{n}{n-m} \Big)^{1/lpha} rac{L(n)}{L(n-m)} + 1 \Big\}. \end{aligned}$$

• For every pair (m, n) of integers, with  $1 \le m < n$ ,

$$\begin{split} b_m b_n \bigg| P(S_m = \kappa_m, S_n = \kappa_n) - P(S_m = \kappa_m) P(S_n = \kappa_n) \bigg| \\ &\leq C \cdot L(n) \bigg\{ n^{1/\alpha} \Big( \frac{1}{e^{(n-m)c}} + \frac{1}{e^{nc}} \Big) + \frac{\frac{m}{n}}{\left(1 - \frac{m}{n}\right)^{1+1/\alpha}} + \\ &+ \bigg( \frac{\frac{m}{n}}{\left(1 - \frac{m}{2n}\right)^2} \bigg)^{1/\alpha} \bigg\}. \end{split}$$

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The Local Limit Theorem and the Almost Sure Local Limit Theo

#### Theorem

Let  $(Z_n)_{n \ge 1}$  be a sequence of centered random variables with finite variance. Suppose that there exists a constant  $\beta > 0$  such that, for all integers  $m \ge 0$ , n > 0,

$$E\left[\left(\sum_{i=m+1}^{m+n} Z_i\right)^2\right] \le C\left((m+n)^\beta - m^\beta\right),\tag{4}$$

for a suitable constant C independent of m and n. Then, for each  $\delta > 0$ ,

$$\sum_{i=1}^{n} Z_i = O(n^{\beta/2} (\log n)^{2+\delta}), \quad P - a.s.$$

#### Theorem

Let  $\alpha > 1$  and assume that there exists  $\gamma \in (0,2)$  such that

$$\sum_{k=a}^{b} \frac{L(k)}{k} \leq C(\log^{\gamma} b - \log^{\gamma} a).$$

Then  $(X_n)_{n \ge 1}$  satisfies an ASLLT, i.e.

$$\lim_{N\to\infty}\frac{1}{\log N}\sum_{n=1}^{N}\frac{b_n}{n}\mathbf{1}_{\{S_n=\kappa_n\}}=\Lambda g(\kappa).$$

## Main steps

We apply the above theorem to the sequence

$$Z_j := \sum_{h=\rho^{j-1}}^{\rho^j-1} \frac{Y_h}{h}.$$

$$Y_n = b_n \Big( \mathbb{1}_{\{S_n = \kappa_n\}} - P(S_n = \kappa_n) \Big).$$

We obtain that

$$\frac{1}{n}\sum_{1\leqslant h\leqslant \rho^n}\frac{Y_h}{h}\to 0,$$

whence (as before)

$$\frac{1}{\log n}\sum_{1\leqslant h\leqslant n}\frac{Y_h}{h}\to 0.$$

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